# More than a She-recession: Long-term feminization and short-term pandemic effects 

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#### Abstract

The Covid-19 crisis has been defined as a "She-recession" because of its disproportionate impact on female employment by contrast to past recessions defined as "Man-recessions", for the usual disproportionate impact on men. The roots of the Sherecession can be however traced back to the persistence of gender asymmetries both intra-household and extra-household in the labour market, a phenomenon known as feminization. This paper aims at measuring and explaining the gender differences in the impact of the Covid-19 crisis on the Italian labour market from a macroeconomic perspective. We measure the duration, depth and diffusion of the Covid-19 crisis on job losses, structural unemployment and inactivity. We find that the impact of the Covid19 crisis has been more than proportional for women, especially for low educated female workers and working in the South during 2020.


Keywords: feminization, hysteresis, labour markets<br>JEL Classification: J16, E32, J6

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## 1 Introduction

During the pandemic crisis, a new concept has spurred into both academic and journalistic literature meant at qualifying from a gender perspective the recessionary impact of the crisis on labour markets, namely, the notion of She-recession. The very concept of She-recession tries to account for the fact that a crisis might have an asymmetric unfolding with respect to its gendered impacts and takes origin from the comparison between the Great Recession in the US, from 2007 to 2009, vis-à-vis the Covid-19 crisis in 2020. In fact, the literature defines the past economic downturns as "Man-recessions", since the impacts of pre-pandemic crises have usually been harder for male employment rather than for female one. Until the pandemic crisis, economic shocks have been mainly affecting male-prevalent industrial sectors, as manufacturing and construction, subject to a high cyclical volatility (Hoynes et al. (2012); Rubery and Rafferty (2013); Alon et al. (2021)). Differently, the Covid-19 crisis has recorded a major decline in female employment (Shibata (2020)). The origin of the crisis, a global health crisis, and the adopted measures to contain it, in particular social distancing and lockdowns, have largely impacted upon service sectors, the latter involving working activities characterised by inter-personal contacts and which can not be easily executed from home. The sectors are worldwide characterised by predominant shares in female employment. In addition to closures of workplaces, school closures have harshly impacted on motherhood, with female workers having to split their time between homeworking, whenever possible, and childcare, which typically weights on women's shoulders. The interaction between the specificity of the pandemic crisis and the ensuing non-medical containment measures has resulted in what has been now commonly understood as She-recession.

According to Del Boca et al. (2020), Italy is an interesting case study on this matter. On the one hand, it is the first European country where the coronavirus has spread and where very strict lockdown measures have been adopted, particularly long school closures. On the other hand, the Italian labour market is characterised by low female labour force participation and long-lasting gender asymmetries. Cetrulo et al. (2022) show that Italian women are mainly employed in essential, low-skilled sectors, service and retail activities, which could not have been executed from home during lockdown phases; largely under temporary contracts or self-employed, many women were not covered by the firing restrictions applied by the Italian government as a response to the Covid-19 related economic crisis. Taken together gender asymmetries in the sectoral distribution of occupations and the precarious
working conditions, female workers in the Italian labour markets have been dramatically exposed to negative side effects of the pandemic, with peaks in unemployment and transition to inactivity at $98 \%$ during the last quarter of $2020^{1}$. The roots of the Sherecession can be however traced back to the persistence of gender asymmetries both intra-household and extra-household in the labour market, a phenomenon known as feminization. ${ }^{2}$ With extra-household feminization we intend feminization of the labour market implying occupational segregation in low-paid activities, contractual segregation in temporary and fixed contracts, fragmentation of jobs. Intra-household feminization implies gendered division of unpaid and care work and gendered norms, the latter reflected into high inactivity and low labour market female participation.

In this contribution we ask the following research questions: has the She-recession taking place in Italy as a consequence of the Covid-19 shock? If yes, how can we measure the severity of such phenomenon? To what extent the root causes of the She-recession are temporary or, alternatively, structural? We try to answer to these questions by studying the unfolding over time of the She-recession in Italy. We take a macroeconomic-structuralist perspective, identifying its root-causes in the persistent attributes of female employment conditions in the Italian labour market, both in the long and in the short run. We first present a series of structural labour market trends in female employment. We then assess the She-recession looking at depth, duration and diffusion of the Covid-19 crisis, according to the NBER methodology of crisis identification, considering the long-term pre-crisis trends.

[^1]In this respect, we build and refine the statistic developed by Fazzari and Needler (2021), a loss function measuring the unfolding of the crisis in terms of the three above mentioned dimensions. The measure represents a proxy to detect the potential emergence of hysteresis in labour market patterns, being built upon deviation from long-term trends.

Using quarterly data from the Italian Labour Force Survey, we compute losses functions in employment, structural unemployment and inactivity distinguishing by gender, education level, geographical location and sectoral distribution. The measure allows to retrieve equal vs unequal outcomes in the distribution of losses among categories for each series of interest. We document a disproportionate impact on female employment and a proportional effect on female inactivity due to previous hysteresis in the labour market. Women from the South and women with lower education levels have been the most exposed categories, recording an impact more than proportional with respect to higher educated women and women from other regions.

The analysis is structured as follows: Section 2 accounts for a theoretical discussion and previous findings on She-recession; Section 3 looks at structural trends and changes in female labour markets, focusing on the role of education, tertiarization, flexibilization of the labour market, and regional asymmetries, in a nutshell, patterns of feminization of labour markets (Manicardi 2023, this issue). Section 4 presents the data, Section 5 links structural trends and pandemic effects in female labour markets to detect the Sherecession and present the methodology and the results. Our conclusions are laid out in Section 6.

## 2 She-recession: why the pandemic crisis is different

### 2.1 Female labour demand and supply during recessionary periods

In order to understand the specific attributes of the She-recession we need to compare such an event with previous episodes of crises. Event crises are generally understood as being originated from supply or alternatively demand "shocks". Generally speaking, with respect to supply shock crises, the literature refers to energy crises or imported-inflation of intermediate goods that propagate via price-channels, with the oil crisis in the seventies being the textbook case. When coming to demand crises, less acknowledged till the Great Recession in 2008, the literature refers to declines in wage growth and lack of consumption. The pandemic crisis has been classified as both a supply crisis, due to the reduction of
working hours and labour supply, and a demand crisis, as far as wage compression and income losses might have constrained demand.

Crises, particularly in labour markets, propagate via industry channels: working hours reduction and employment losses derive from output contraction in industries and sectors more exposed to the shocks. Therefore, both industry composition in terms of gender, and country composition in terms of industry, determine the severity of each given crisis and any eventual disproportionate effect on specific labour market/industry segments. Comparing the pandemic with "ordinary" shocks provides evidence of asymmetric effects in terms of gendered job losses.

Hoynes et al. (2012) compare the financial crisis of 2007 with the economic recession in the 1980s in the United States. The loss in overall employment has been stronger and longer during the Great Recession than in the 1980s recession, mainly because of a higher participation of women to the labour market, but men accounted for the highest share of job loss and they find that unemployment was more cyclical for men than for women both during the 80s and during the financial crisis. Rubery and Rafferty (2013) trace back gender differences in crises outcomes to job segregation, work flexibility and participation, and welfare state capacity. The latter are channels through which there could be propagation of negative shocks leading to unexpected outcomes. Occupational segregation can both protect or expose more women according to the characteristics of the recession and, because of the low pays they receive, they might record an increase in employment during recoveries, acting as a substitute for higher paid male workers; in addition, female labour supply is hugely dependent from the household role and on the welfare state benefits in kind. These attributes make the female component of labour markets quite distinctively behaving vis-à-vis male ones. In line with this expectation, although highly shock-exposed industries have usually been male-dominated ones, looking at cyclical volatility between the last quarter of 2019 and the second one of 2020, female predominant industries have recorded the highest loss in employment (Alon et al. (2021)).

Therefore, the very nature of the recession in itself is the first reason of gendered effects of crises. However, in addition to asymmetric exposure to shocks because of occupational segregation, resulting from asymmetric labour demand patterns, also female labour supply patterns are quite specific.

Female responses to recessions might be very diverse. During ordinary shocks, a counter-
cyclical response of female employment to counterbalance the risk of unemployment of the male partner has been defined by Lundberg (1985) as the "added-worker effect" acting as an insurance against employment uncertainty, credit constraints and loss in earnings of the male members of the household. Rubery and Rafferty (2013) argue that in past recessions before the Great Recession in 2007, cyclical volatility of female employment was found only in manufacturing, where women mainly occupied buffer positions, with flexible hour arrangements in order to adjust to variations in demand and protect the male positions at the core, while in other sectors they were more protected. Flexible participation to the labour market instead has led to an increase in unemployment and a decrease in inactivity rate for women between 2007 and 2011. The evidence of the added worker effect varies across countries. Bredtmann et al. (2018) analyse the responsiveness of female labour supply as consequence of male partners' unemployment status across different welfare regimes. They find that only Mediterranean countries (Italy, Spain, Portugal and Greece) show an added worker effect both on the extensive and intensive margin, because of low social protection and dominance of the familistic culture. Similarly, Prieto-Rodrıguez and Rodrıguez-Gutiérrez (2003) find the added worker effect to be present in the years 1994, 1995 and 1996 only in Italy with respect to other countries in their sample, where female labour supply is less related to the employment status of their husbands. Muñoz et al. (2015) find the presence of the added worker effect during the Great Recession in Italian and Spanish households. According to the authors, the added worker effect is more significant for women while labour supply decreases for mothers and Italian women when income in the household increases. Ghignoni and Verashchagina (2016) find evidence for the added worker effect in Italy during the Great Recession, the probability to participate to the labour market to be $4.9 \%$ higher for women whose husband has become unemployed during 2010-2012. Childcare weaken the added worker effect, while mortgages positively affect the increase of female participation early in the crisis (2006-2008). However, employment translates mainly in low pay and precarious jobs. Franceschi (2011) estimates the added worker effect in Italy to be $2.4 \%$ and $4.2 \%$ for employment and participation respectively in 2009-2013 and that the added worker effect accounts for $8 \%$ of the increase in female employment in 2012 and 2013. The effect is lower for mothers while there is no evidence of male "added worker effect" as a consequence of job loss for women in the household. Baldini et al. (2018) analyse the reaction of households members to an employment shock, confirming the results of Franceschi (2011) and specify
that for households in the South during the recession labour supply was intensified.
Remarkably, the Covid-19 crisis did not show these regularities. Because of the magnitude of employment decline during the pandemic (much higher than for the Great Recession, Shibata (2020)), the "discouraged worker effect" (Lundberg (1985)) characterised both female and male employment dynamics. Indeed, according to Albanesi and Kim (2021a), in the US employment decreases by $8.5 \%$ for married men with children vis-à-vis $13 \%$ for married women with children during the Covid-19 crisis. While the impact on male employment is higher with respect to the Great Recession ( $-2.4 \%$ versus $-8.5 \%$ ), the relationship is reversed for women ( $-13 \%$ in the pandemic crisis versus $+0.2 \%$ during the Great Recession). According to Kim et al. (2022), less educated women, non-married and non-married with children are the "she" of the Sherecession, while the "he" are never married and low educated men.

Overall, the impact of the Covid-19 crisis on female employment has been twofold. On the demand side, acting via the industry composition channel, social distance measures have impacted female workers more than male ones due to the gender composition of the industries more subject to closures. On the supply side, acting via the childcare channel, school closures and gender norms in child and elder care forced working women to reduce their supply of labour also in sectors less hit by the pandemic, with the possibility of remote working, or even to exit the labour market.

### 2.2 She-recessions across countries

The extant literature shows that there is a large cross-countries variation in the impact of the Covid-19 crisis on female employment. Alon et al. (2021) find that (i) gender gaps in employment effects decrease when controlling for occupations, but gender gaps for parents with school-age children are the largest in the United States and in Canada; (ii) the effect on employment is small in Germany and in the Netherlands but high on working hours, particularly because of furloughs schemes preserving employment but reducing working hours especially for women; (iii) the relative worked hour index for Swedish workers, that is the ratio between worked hours by women and worked hours by men between the last quarter of 2019 and the second quarter of 2020, has increased given that Sweden has adopted softened school closures; (iv) in Spain the industry channel has a decisive impact since the results show a significant decline in hours for women with school-age children when controlling for
occupation types, while it was not significant in the regression with no industry channel effects.

Hupkau and Petrongolo (2020) find that the effect of the Covid-19 crisis has been mixed in the United Kingdom: on the one hand, the gender gap has narrowed since the impact of the recession has been equally distributed on men and women in terms of job losses and resort to furloughs, while on the intensive margin, the reduction in worked hours and in earnings have been even smaller for women and an increase in fathers devoting time to housework and childcare is observed; on the other hand, women have taken care of $60 \%$ of the additional childcare hence gender differences in family life have widened. By the same token, for Italy Del Boca et al. (2020) and Del Boca et al. (2021) find that the time devoted by women for childcare is always higher except for those women who kept working where they used to also during the pandemic. At the same time, they find that women's working arrangement does not affect the time their male partners devote to childcare or housework, but males' working arrangement does on their female partners. Biroli et al. (2021) run a similar analysis comparing Italy, the United Kingdom and the United States while Djankov et al. (2020) find that the gender gap in labour force participation has shrunk in Denmark, Norway, Australia and the United Kingdom thanks to ad-hoc policies.

Particularly for Italy, the distribution of occupations by gender together with the housework division of labour are the main channels for which the impact of the pandemic has been so harsh for women. Cetrulo et al. (2022) show that women are mainly employed in essential, low-skilled sectors, such as service and retail activities, largely non teleworkable occupations, with temporary contracts, or self-employed. The pandemic has exerted two consequences on these workers: firstly, they were not covered by the firings' restrictions, thus they were the category of workers for which the highest job losses were recorded; secondly, the impact of the lockdown measures on these sectors has exacerbated the polarization in wages between precarious and protected workers, widening gender gaps. Indeed, only $30 \%$ of professions can be executed from home and women represent a low share of them (mainly employed in the administrative sector).

## 3 Beyond recessions: Long term patterns of feminization of labour markets

The supply-side responses of women to the Covid-19 crisis acknowledged by the literature are particularly driven by gender norms. Fabrizio et al. (2021) find that women without children whose occupation implied inter-personal contact have recovered in terms of employment much faster than mothers having similar occupations. Other than gender association for women in their role of mothers, the gender gap in earnings justifies the choice to reduce working hours for women who usually earn a lower salary than their male partners (Kleven et al. (2019); Albanesi and Kim (2021a); Albanesi and Kim (2021b); Hupkau and Petrongolo (2020), Sun and Russell (2021)). However, the root causes of such patterns have to be traced back to long term sources of occupational segregation.

The combination of structural changes in job opportunities in new growing sectors and higher level of education have constituted a turning point for female employment. During the seventies, new job opportunities for women were constituted by the access to previously typical male professions as teachers and bank clerks, by the increase in demand for health and child care, a sector that already was female-dominated, and by technological change particularly in clerical work and telecommunications (Blau et al. (1998)). Technological change, mainly computerisation and the telecommunications revolution in the 1970s, has induced a recomposition effect in tasks and functions executed in the workplace, partly covered by female administrative jobs, growing also as a consequence of an increase in firm size and the need for clerical work (Costa (2000)). Newly demanded job tasks matched with women's abilities, thanks to an increase in their access to education (Goldin (1984); Dolado et al. (2002)). The increase in women accessing higher levels of education was a result of what Goldin (2006) defines the quite and revolutionary phase of the modern economic role of women. Starting in the late seventies, women's identity of themselves and decision making have started to shift from gender association, for which women identified themselves as housewives not participating to the labour market or not investing in their education, to own individuality, involving a stronger attitude towards economic decisions, participation to the labour market and investments in human capital and career development in a long run horizon.

In the nineties, the tertiarization of the economy drives female employment, providing


Figure 1: Employment and participation rates by gender between 1977 and 2020 in Italy, age 15-64. Data Source: Labour Force Survey, ISTAT
for new job opportunities in services. Patterns in Italy follow such a trend. In Figure 1, both participation and employment rates are characterised by a positive and increasing trend for women, starting from levels below $40 \%$ for both rates in the 1977 and reaching the $50 \%$ - $55 \%$ in 2020. Nonetheless, both rates are still low compared to EU level averages ${ }^{3}$, and especially lower than male participation and employment rates.

Despite the increasing trend in female employment, gender norms keep affecting the decision to participate to the labour market, the level and duration of education and consequently the final occupation and sector of belonging for women, leading both to concentration of female employment in low-value added sectors and/or limited access to high professional status (horizontal and vertical occupational segregation). Employment discontinuity in particular is among the sources of segregation (Goldin (1984)). According to

[^2]Goldin (1984) women, as mothers and wives, exiting and entering again later the labour market do not become doctors or executives but nurses and typers (Goldin (1984), p.2).

Although the increase in female education has narrowed the gap with male education level (Altonji and Blank (1999)), there is a wide agreement in the literature about the fact that the improvement in females' job experience given by non-intermittent careers, tenureship and training accounts more than education in narrowing differences in earnings (O'Neill and Polachek (1993), Wellington (1993), Blau and Kahn (1997), Goldin (2006)). Interacting discontinuity in employment and education choices, Mincer and Polachek (1974) argue that there is a positive relationship between the continuity in participation, mainly observed in single, young women without children, and the initial investment in human capital in the first working experience, more than in later ones, while because of discontinuity of participation both women, usually married with children, and employees do not invest in learning and training skills.

Employment discontinuity and hiring in discrimination have long-lasting root causes. Historically, in Italy between 1951 and 1971 the female participation rate has decreased with the shift from land to factory work (Costa (2000)). In that period exclusively young, non-married women and with no children were employed in factories, mainly with parttime contracts and non-regulated dismissals. At the same time, care work and "industrial homework" were performed by up to 1 million women, in particular those migrating from the South to the North of Italy, in precarious work conditions in the informal economy. Such pattern is observed during the 1970s as well, as a consequence of the feminization of industrial homework and exacerbated in the 1980s due to the flexibilization of the labour market (Betti (2016)).

Employers' perception of motherhood as a source of potential absenteeism from work, lower commitment and lower productivity affects female employment status from the hiring process. Acker (1990) claims that organizations are gender-structured, according to which a job, a set of tasks, competencies and responsibilities in the organization's hierarchy, need to be filled by a worker with no other responsibilities outside the ones defined within the job's boundaries. As a consequence, a woman can't fit the "abstract job" since women have obligations in their household life other than those required by the job, while men have not. Cultural perspectives driven by gender norms and statistical discrimination (employers who rely on group averages based on stereotypes about candidates' human capital which can't
be standardised from a job application (Birkelund et al. (2022), González et al. (2019)) are identified by the literature as the main drivers of gender discrimination by employers in hiring. Stereotypes can be descriptive, for which men are more committed to work and have leadership qualities, while women are caregivers. Descriptive stereotypes generate prescriptive ones which establish the women's role as "good mothers" and men as "ideal workers" for which women are expected to be less committed to work, more absent and thus "less productive" fuelling occupational segregation into male and female dominated sectors (González et al. (2019).

The above mentioned channels, namely, employment discontinuity and intermittent careers, hiring discrimination because of motherhood, job-title discrimination because of different "innate" abilities between men and women, are all factors behind what we term long-term feminization of the labour market. In the next paragraphs, we outline trends in Italian female employment with respect to education, sectors, contracts and geographical distribution, to provide some supportive statistical evidence of such long-term trend. ${ }^{4}$

Figure 2 shows the female participation and employment rates by level of education between 1977 and 2020 in Italy: the higher the level of education, the higher the participation rate. Figure 3 presents the shares by education level within female employment. The improvement in education is straightforward: in 1977, only $4.4 \%$ of employed women had tertiary education level while in 2020 the share is $31.2 \%$; the share of working women with a upper secondary level of education follows a similar trend and in 2020 accounts for the highest share of employed women (46.2\%). In 1977, primary education level accounted for the category with the highest share in employment, while the share of working women with lower secondary education level is quite stable over time.

Figure 4 provides for employment shares by gender in macro-sectors, namely agriculture, industry and services in 1977 and in $2020 .{ }^{5}$ Female employment is one third of male one in agriculture in 2020, a stable one fourth share in industry, one third in services in 1977 and almost equal to male employment in 2020. Indeed, with the tertiarization of the economy,

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Figure 2: Shares of female participation rate by education from 1977 to 2020, age 15-64. Data Source: Labour Force Survey, ISTAT


Figure 3: Female employment rate by education from 1977 to 2020, age 15-64. Data Source: Labour Force Survey, ISTAT
women were employed mainly in administration services and public sectors (Betti (2016)). We then look at the percentages of employment by gender in the different service activities, being the latter a female-dominated macro-sector. From Figure 5, adopting a NACE Rev. 2 classification, "Education and health" is the sector with the highest share in female employment ( $72 \%$ vis-à-vis $28 \%$ of male employment both in 2008 and in 2020) and "Other service activities" ( $68 \%$ via-à-vis $32 \%$ for male employment in 2020). "Accommodation and food" and "Administrative and support to firms" shares are quite balanced, while for "Information and communication" and "Transportation and storage" male employment is much higher. This evidence accounts for the pattern of horizontal female occupational segregation, largely concentrated in so called low value added activities. Looking at occupational distribution by professional categories, a proxy for vertical segregation, Figure 6 shows that only $33 \%$ of executives are women in $2020,45 \%$ middle managers and $31 \%$ self-employed, while they account for $57 \%$ of white collars. Shares in professionals, middle management and executives increase between 2004 and 2020. However, the bulk of female jobs being concentrated in "Other services" and in "Education and Health" are less subject to vertical mobility. In fact, among sectors, equally defined occupations are not equally remunerated, and horizontal segregation is not neutral and exacerbates vertical disparities. Typically, in low value-added sectors, where female employment is concentrated, the possibility of professional upgrading and vertical mobility are lower than in so called high value-added sectors. Therefore, any potential amelioration in vertical mobility has to compared with stagnant horizontal segregation.

Figure 7 shows that women account for the lowest share in full time and open ended contracts both in 2004 and in 2020. The flexibilization of work since the 1990s has become a feminized pattern especially in Italy where both low and high skilled women are employed in occupations mainly by atypical contracts (Betti (2016), Manicardi (this issue)). According to Petrongolo (2004), the high shares of part-time contracts for women in Southern Europe is due to the gender discrimination affecting such countries rather than for voluntary women choices as in Northern countries, given that sectors as hospitality, cleaning, education, care, personal and social services happen to be both characterised by part-time/temporary working arrangements and female dominated, sectors in which women are considered as "more appropriate" (Buckingham et al. (2021), Birkelund et al. (2022)).

Italy is characterised by a persistent asymmetry between Northern and Southern regions


Figure 4: Percentage share in employment by gender in macrosectors in 1977 and in 2020 in Italy, age 15 and over. Data Source: Labour Force Survey, ISTAT


Figure 5: Percentage share in employment by gender in the service sectors in Italy in 2008 and 2020, age 15 and over. Data Source: Labour Force Survey, ISTAT.


Figure 6: Percentage shares in employment by gender and by professional status in 2004 and in 2020, age 15 and over. Data Source: Labour Force Survey, ISTAT


Figure 7: Percentage shares of full/part time and fixed/open-ended contracts in employment by gender in 2004 and 2020, age 15-64. Data Source: Labour Force Survey, ISTAT
in gender labour market outcomes. Among the many territorial divides, gender divides are quite striking. Figure 8 reports the female structural unemployment rate (unemployed status for more than 12 months) ${ }^{6}$ and the inactivity rate by region. Structural unemployment is higher more than twice for women in the South with respect to more virtuous regions in the Centre and in the North of Italy. The inactivity rate is even more alarming suggesting strong discouragement effects in female labour market participation in half of the country. Notably, the South is also characterised by strong familistic orientations of the division of labour, with individuals pursuing the traditional division of labour following the male bread-winner model, with gender norms, at the very least, discouraging female participation in formal labour markets and women emancipation from family ties. In addition, from the demand side of the labour market, the structural weakness of the southern productive structure hampers employment opportunities, and particularly good ones.

The documented patterns highlight the process of feminization of the labour market since the seventies. Such process can be expressed by the following trends: (i) female occupational segregation in low value-added sectors, (ii) unmatched increases of educational attainments with professional upgrading and in general wage remuneration for female workers, (iii) female disproportionate exposure to unstable and flexible contractual regulations, (iv) persistence of gender norms reproducing intra-household asymmetric division of labour and extra-household labour market participation.

## 4 Data

Having documented such long term patterns, we now intend to focus on the specific effects on female workers due to the pandemic unfolding. In particular, we are interested in capturing hysteresis effects due to long-run trends in feminization as potential amplifiers of the pandemic crisis.

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Figure 8: Female structural unemployment and inactivity rate by regions in Italy in 2020. Data Source: Labour Force Survey, ISTAT.

In order to accomplish the task, we analyse quarterly data from the Labour Force Survey of the Italian National Institute of Statistics (ISTAT) of employment, structural unemployment, inactivity for women and men between 15 and over years old ${ }^{7}$ accounting for a long term period, from 1993 to 2020 and of female employment by macro-regions (North, Centre and South, islands included) and by level of education (tertiary, upper secondary, lower secondary and primary) for women with age 15 and over from 2000 to 2020 and by sectoral distribution (macro-sectors and focus on 9 service sub-sectors, NACE-1 digit) for women with age 15 and over from 2008 to 2020. Sectoral data are seasonally adjusted. Tables C.1, C.2,C.3, C. 4 and C. 5 in the Appendix show the main descriptive statistics about these key variables. t is the number of observations namely the number of quarters of the time series that changes because of data availability. On average over the time period under consideration, female employment is lower than male employment by $4,572,650$ units, female structural unemployment is higher than male one by 2,000 units while female inactivity is higher by more than 6 million units. Looking at the mean values of female employment by regions, employment in the North almost doubles employment in the Centre and the South, while women with upper secondary education level show the highest mean value across the educational distribution. Women are mainly employed in the service sectors, accounting for almost 8 million workers on average. The sub-sector employing the highest number of workers is education, human health and social work activities with more than 2.5 million workers on average (the so called care sector), wholesale retail trade, reparation of motorvehicles and motorcycles with 1.3 million workers, administration and support services and other services over 1 million workers on average along the 2008-2020 period under consideration.

## 5 Detecting She-recession

### 5.1 From long-term feminization to hysteresis

Economic shocks, as the Covid-19 crisis, may induce hysteresis effects, impacting structural unemployment in particular. Dosi et al. (2018) identify three main channels of recessions

[^5]inducing hysteresis in unemployment and output growth: decrease in productivity, skills deterioration and in entry dynamics. In relation to unemployment, during recessions firms fire workers and unemployment rises. If recessions are deep, recovery will be slower and less powerful, causing an increase in the duration of unemployment which, in turns, implies skills deterioration: long-term unemployed workers stop learning by doing processes, miss to acquire new techniques of production, are less probable to find a job, and whenever it occurred, their skills are deteriorated and their productivity will be lower. Two important side effects emerge. First of all, an increase in the duration of unemployment raises structural unemployment, typically hysteretic via skills' deterioration which induces further lengthening of unemployment duration. At the same time, structural unemployment largely characterises female labour market status. In Italy, female structural unemployment, with a duration longer than 12 months, has always been higher than the male one, and, on the contrary, employment and participation rates have always been much lower for women than for men $^{8}$, and below EU average. ${ }^{9}$

As a consequence, long-term unemployed people may exit the labour market, because of the strong feeling of discouragement and transiting into inactivity. Discouragement has particularly characterized the labour market dynamics during the Covid-19 crisis both for men and women: ISTAT counts a trend variation of 1 million and 200 units of inactive workers in March 2020 with respect to March $2019^{10}$. The increase in inactive workers during the pandemic can be interpreted as a consequence of the persistence of precariousness and high unemployment in the Italian labour market.

In fact, occupational segregation, access to education and precariousness of female labour market participation present strong persistence over the long run, and in that have affected the labour market dynamics during the pandemic phase. In Figure 9, the autocorrelation functions for structural unemployment, employment and inactivity by gender are shown.

[^6]All variables, both male and female, are strongly correlated with their past values, although with a decaying memory. The time lasting memory of the variables provides for a first evidence about how individual current status in employment, unemployment, inactivity is deeply affected by past outcomes.

Figure 10 shows the change in the employment, structural unemployment and inactivity rate in total and by gender from the last quarter of 2019 to the last quarter of 2020 . The strong increase in the inactivity rate between the first and the second quarter of 2020 is evident both for men and women. In parallel, both the employment and the unemployment rate decrease, the latter being higher for women than for men. These trends give evidence of the strong discouragement effect impacting the Italian workers during the Covid-19 crisis.

Figure 11 shows the level in employment by gender during 2020 and the percentage changes from quarter to quarter. The decrease is stronger for women only during the second quarter of $2020(-2.24 \%$ vis-à-vis $-1.04 \%$ for men), despite male employment increases during the third quarter, it decreases during the last one while female employment increases.

Figure 12 shows the percentage change from previous quarter in employment by gender in macro sectors. Looking at the impact on total employment in the first quarter of 2020, it is evident the decline in all sectors with respect to the last quarter of 2019, as in the second quarter of 2020 in particular for wholesale, retail and trade, accommodation and food and services. The decrease is harsher for female employment. In the third quarter of 2020, both female and male employment in wholesale, retail and trade, accommodation and food and in services overall increase with respect to the previous quarter, while manufacturing and other services (including arts, entertainment and recreation) keep decreasing. It is interesting to notice the relevant changes in female employment in constructions and agriculture. As claimed by Rubery and Rafferty (2013), women cover buffer positions in male dominated sectors thus are more subject to cyclicality.

Within the service sector (Figure 13), accommodation and food, wholesale and retail trade, other services arts, entertainment and recreation, other service activities ${ }^{11}$ have been the most affected. Male employment in accommodation and food, as expected, shows the

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Figure 9: Autocorrelation in the Italian Labour market by gender, quarter lags, age 15 and over. Data Source: Labour Force Survey, ISTAT.


Figure 10: Employment, structural unemployment and inactivity rates by gender in Italy during 2020, age 15 and over, quarterly data. Data source: Labour Force Survey, ISTAT


Figure 11: Levels of employment and quarterly changes by gender in Italy during 2020, age 15 and over. Data Source: Labour Force Survey, ISTAT


Figure 12: Employment changes from previous quarter in macro-sectors by gender during 2020, age 15 and over. Data Source: Labour Force Survey, ISTAT
highest variability across quarters, with a seasonal increase in the third quarter then followed by a decrease in the fourth quarter of 2020. For women the losses look to be more structural and persistent: while the summer increase in accommodation and food is much lower when compared to the decrease in the final quarter, the losses in employment in other sectors are quite persistent even during the final quarter, except for the care one. Compared to men, female employment shows deeper negative changes in transportation and storage and in wholesale and retail trade during the first quarter of 2020, in wholesale and retail trade in the second quarter and in other service sectors as financial, communication activities, other services, public administration in the last quarter of 2020.

While exhibiting long-term memory patterns, typical of unit root variables, the analysis of autocorrelation functions is not enough to actually measure persistence in duration of negative shocks for labour market outcomes. In the next section, we are going to focus on the measurement of hysteresis in labour market outcomes for female workers and the extent to which a She-recession has occurred.


Figure 13: Employment changes from previous quarter in the service sectors by gender during 2020, age 15 and over. Data Source: Labour Force Survey, ISTAT

### 5.2 Loss functions and their distribution

The evidence presented so far, although accounting for structural patterns and pandemic labour market effects, is still not conclusive of the extent to which a She-recession has occurred. We advance therefore with a direct measurement of losses in line with the methodology proposed by Fazzari and Needler (2021). The indicator measures whether the impact of the Covid-19 crisis has been disproportionate for a given gender category. We compute such statistics for employment losses, structural unemployment and inactivity with respect to the pre-recession trends. To account for the severity of the event in a hysteresis type of framework (i) we select structural unemployment (unemployment longer than 12 months, not simple unemployment) and inactivity as variables of interest being characterised by strong persistence; (ii) our statistics, being loss functions, assess the depth, duration and diffusion of the Covid-19 crisis, measuring deviations and duration of such deviations from the pre-crisis trend. In addition, focusing on the female employment impact, we then analyse heterogeneity across levels of education, regional and sectoral distribution within female employment to identify their underlying role in the disproportionality/proportionality of the pandemic effects.

We improve the methodology of Fazzari and Needler (2021) by (i) capturing hysteresis on the Italian labour market by considering structural unemployment and inactivity other than
employment; by (ii) adopting a long run perspective filtering the data with a trend starting in 1993 in order to account for long-run deviations from existing trends, while Fazzari and Needler (2021) apply a short run trend; (iii) by adopting the Christiano-Fitzgerald filter instead of a linear trend. The Hodrick Prescott filter is applied as a robustness check to show that the results are not driven by the choice of the filter. The use of the filter allows to clean from the short-run components of the series and to focus on the structural ones, being our interest on the lasting impact of the feminization of the labour market as possible pre-condition explaining the effects of the short-run Covid-19 crisis.

In order to compute the key statistic, we follow six main steps:

1. prior-recession peak and trough identification. We firstly define the recession period following the three main criteria adopted by the NBER ${ }^{12}$ : depth, duration and diffusion.

As cited by Claessens et al. (2009), "a recession begins just after the economy reaches a peak of participation and ends as the economy reaches its trough". In terms of depth, a recession can be defined as severe when output declines more than $3.15 \%$ (Claessens et al. (2009)); its duration is defined between the first peak and the first trough, namely by the time span between the first decrease of $-3.15 \%$ (or more) and the first increase grater or equal to $+3.15 \%$. To address depth and duration, we identify the prior-recession peak and the first trough looking at GDP growth, following NBER Business Cycle Dating Committee guidelines.

Figure 14a shows the annual trend variation of GDP growth (grey line) and of total employment (dashed line) between the last quarter of 2019 and the second quarter of 2021 with respect to the correspondent quarter of the previous year. In addition, we also show the cumulative variation of GDP in order to give an account of the persistent negative dynamics of the variable. The prior-recession peak is the last quarter of 2019 since the first drop is lower than $-3.15 \%(-6.36 \%)$ and occurs during the second quarter of 2020. The trough is the last quarter of 2020 , since the first trend variation over $+3.15 \%$ is the first quarter of 2021 . The end of the recession is then the last quarter of 2020 (2020-Q4). Total employment follows the annual trend variation of GDP.

[^8]Figure 14b shows the cyclical variation of GDP growth (grey line) and employment (dashed line), namely the percentage change with respect to the previous quarter, together with the cumulative GDP variation (red bars). Despite the cumulative cyclical variation is still negative, the GDP growth increases by about $14 \%$ during the third quarter of 2020, thus the trough is the second quarter of 2020 (2020-Q2).

Therefore, we identify the last quarter of 2019 as the pre-recession peak and the last quarter of 2020 as trough with respect to the trend, while the second quarter of 2020 with respect to the cyclical variation. Estimates are pursued using both time spans. As a result, (i) the identification of the peak and the trough allows to measure the duration of the recession, to define the time span of the impact of the Covid-19 crisis on the Italian labour market (longer for annual trend with respect to cyclical variations); (ii) the identification of the peak and the trough with respect to the severity of GDP growth variation explains the depth of the recession, while actual data of the variables of interest at the peak signal the proportionality of the impact for the category of worker we are considering; (iii) comparing employment, structural unemployment, inactivity and female employment by education, by regions and by sectors explains the diffusion of the recession.
2. Isolation of trend and cyclic data components. To capture the impact of the Covid-19 crisis, we look at the differences between actual and filtered data over the defined recession period. In particular, the filter captures the state of the economy in the previous phase. However, the trend is calculated over a long-run horizon, starting in 1993 in order to account for any hysteretic patterns in the series analysed, in line with the evidence presented in Section 3.

Each time series is filtered by the Christiano-Fitzegerald filter (CF from now on), to detach the trend from the cycle, namely the variation from trend as a consequence of shocks. The Hodrick-Prescott filter (HP from now on) is applied as a robustness check (description and estimates are provided in the Appendix), since the HP extracts a trend that is comparable to the linear trend adopted by Fazzari and Needler (2021) in our case, so it allows us to make a comparison and robustness checks. ${ }^{13}$

[^9]
(a) Pre-recession peak identification: annual trend variation of GDP growth and employment with respect to correspondent quarter of previous year (from 2019-
Q4 to 2021-Q2 with respect to $2018-\mathrm{Q} 4$ and $2020-\mathrm{Q} 2$ )

(b) Pre-recession peak identification: cyclical variation of GDP growth and employment with respect to previous quarter between 2019-Q4 and 2021-Q2

Figure 14: Annual and cyclical variations in GDP and total employment. Data Source: National Accounts, ISTAT

The CF filter decomposes a time series $\left\{x_{t}\right\}_{t=1}^{T}$ into its trend and cyclical components. Assume we have a stochastic process

$$
\begin{equation*}
x_{t}=y_{t}+\bar{x}_{t} \tag{1}
\end{equation*}
$$

where $y_{t}$ is a process oscillating between $2<p_{l}<p_{u}<\infty$ and the frequencies for which $y_{t}$ has power are limited to $\{(a, b) \bigcup(-a,-b)\} \in(-\pi, \pi)$ where $a=\frac{2 \pi}{p_{u}}$ and $b=\frac{2 \pi}{p_{l}}$. For what concerns our analysis, $p_{l}=6$ and $p_{u}=32$, since cyclical components in a business cycle last from a minimum of six quarters ( 18 months with monthly data, 1.5 years with annual data) and a maximum of 32 ( 96 months, 8 years) hence $a=\frac{2 \pi}{32}$ and $b=\frac{2 \pi}{6}$ (Baxter and King (1999), Christiano and Fitzgerald (2003), Hodrick and Prescott (1997)). $\bar{x}_{t}$ is a process oscillating in the complement interval in $(-\pi ; \pi)$ (Fitzgerald and Christiano (1999)). The CF filter approximate $y_{t}$ with $\hat{y}_{t}$, a filter that is a linear function, a projection of $y_{t}$ onto $x_{t}$ of the raw data $x_{t}$ : for $t=1, \ldots, T$

$$
\begin{equation*}
\bar{y}_{t}=P[y \mid x]=\sum_{j=-f}^{p} \hat{B}_{i}^{p, f} x_{t-j} \tag{2}
\end{equation*}
$$

where $f=T-t$ and $p=t-1$. The weights are chosen to minimise the mean square error between $y_{t}$ and $\hat{y}_{t}$, that is $\hat{B}_{j}^{p, f}$ solves

$$
\begin{equation*}
\min _{\hat{B}_{i}^{p, f}}^{j=-f, \ldots, p} \mid ~ E\left[\left(y_{t}-\hat{y}_{t}\right)^{2} \mid x\right] \tag{3}
\end{equation*}
$$

$x_{t}$ is represented as a moving average of order $q$ to avoid the filter to depend on time and non-stationarity of the series. As a result, we get two time series: a trend and a cycle, representing the deviations from the trend.
3. Loss function definition. We compute a loss function $\mathcal{L}$ comprehending the loss in employment, increase in structural unemployment and inactivity by gender and loss in female regional employment and by level of education and sectoral distribution as
quarterly data, there are no big differences between their method and the HP filter, which is widely used in the literature to analyse economic phenomena with quarterly data (Christiano and Fitzgerald (2003), Baxter and King (1999)). We did not use the Baxter-King filter since it dropped some observations at the beginning and at the end of the sample and given that in our time series the end of the sample corresponds to the Covid-19 crisis we can't drop these observations.
the cumulative sum of the CF cycle values:

$$
\begin{equation*}
\mathcal{L}_{x_{i_{j}}}=\sum_{t=Q 1_{2020}}^{T} c_{x_{t, i_{j}}} \tag{4}
\end{equation*}
$$

where $x=\epsilon, u, i$ is the labour market variable (employment, structural unemployment and inactivity), $i=g, e, r, s$ is the type of individuals we are analysing by gender, female education, female regional employment and female employment across sectors respectively, $j$ is the category for each type namely $j=w, m$ for women or men for $i=g$ gender; $j=p, l s, u s, t$ for primary, lower secondary, upper secondary and tertiary level for $i=e$ female education; $j=n, c, s$ North, South ${ }^{14}$ or Centre of Italy for $i=r$ female regional employment, $j=a, c, i, s$ for $i=s$ female employment by macro-sectors. The analysis is deepened also across 9 service subsectors. $T$ is either the last quarter of $2020^{15}$ or the second quarter of $2020 .{ }^{16}$
4. Share of each category over the total loss function. We compute the percentage of the $\operatorname{loss} \mathcal{L}_{x_{i, j}}$ for each $j$ on total loss function for category $i, \mathcal{L}_{x_{i}}=\sum_{j=1}^{J} \mathcal{L}_{x_{i_{j}}}$

$$
\begin{equation*}
l_{x_{j}}=\frac{\mathcal{L}_{x_{i_{j}}}}{\mathcal{L}_{x_{i}}} \% \tag{5}
\end{equation*}
$$

5. Share of each category of total data at prior-recession peak. We compute the share of each category $j$ of type of individuals $i$ for each variable over time $x_{t}$ of actual data at the prior-recession peak $t=p$, the first quarter of 2019 (2019-Q4):

$$
\begin{equation*}
s_{x_{i_{j}}}=\frac{x_{p, i_{j}}}{x_{p, i}} \% \tag{6}
\end{equation*}
$$

6. Quarter loss (QL) indicator ${ }^{17}$. We take the ratio between the share of losses $l_{x_{j}}$ over the share of actual data of variable 2019-Q4 $s_{x_{i_{j}}}$ :

$$
\begin{equation*}
Q L_{x, i_{j}}=\frac{l_{x_{i_{j}}}}{s_{x_{i_{j}}}} \tag{7}
\end{equation*}
$$

[^10]where the denominator captures the persistence in gender, educational and geographical and sectoral distribution inequality over time. The quarter loss is interpreted as follows:
\[

$$
\begin{cases}Q L_{x, i_{j}} \in[0,1] & \text { impact of the crisis less than proportional }  \tag{8}\\ Q L_{x, i_{j}}>1 & \text { impact of the crisis more than proportional }\end{cases}
$$
\]

### 5.3 Results

In what follows, we provide the results of our estimation of loss functions. Figure 15 shows the female and male losses in employment, structural unemployment and inactivity, corresponding to the area in grey during the recession period between the prior-recession peak (2019-Q4) and the last quarter of 2020 (2020-Q4). Figures B.3, B. 4 and B. 5 in the Appendix show the quarter loss areas with respect to the $2020-\mathrm{Q} 2$ trend. The area in female employment appears wider than for men, while for structural unemployment and inactivity there is no evident gender difference.

(c) Filtered inactivity data by gender by the Christiano-Fitzgerald filter, quarter loss in grey

Figure 15: Employment, structural unemployment and inactivity by gender filtered data up to 2020-Q4, quarter loss in grey

Table 1 reports the quarter loss (QL) indicator and the shares behind it. Column 1

(a) Filtered employment data by gender by the Christiano-Fitzgerald filter, quarter loss in grey.

(b) Filtered structural unemployment data by gender by the ChristianoFitzgerald filter, quarter loss in grey
highlights the trough of the recession period, 2020-Q2 or 2020-Q4, column 2 and 3 show the percentage by gender of the deviation from trend for women and men $l_{i, j}$ where $i=g$ corresponds to gender and $j=w, m w$ for women and $m$ for men, the numerator of the QL indicator; Column 4 and 5 show the share by gender in actual data at the prior-recession peak $s_{i, j}$ which captures the persistence of gender differences in the Italian labour market over time being the denominator of the QL indicator. Column 6 and 7 report the QL indicator for female and male workers respectively.

| Employment |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| T | $l_{\epsilon_{g, f}}$ | $l_{\epsilon_{g, m}}$ | $S_{\epsilon_{g, w}}$ | $s_{\epsilon_{g, m}}$ | $Q L_{w}$ | $Q L_{m}$ |
| 2020-Q2 | 42.905 | 57.095 | 42.157 | 57.843 | 1.018 | 0.987 |
| 2020-Q4 | 57.289 | 42.711 | 42.157 | 57.843 | 1.359 | 0.738 |
| Structural unemployment |  |  |  |  |  |  |
| T | $l_{u_{g, f}}$ | $l_{u_{g, m}}$ | $s_{u_{g, w}}$ | $s_{u_{g, m}}$ | $Q L_{w}$ | $Q L_{m}$ |
| 2020-Q2 | 20.171 | 79.829 | 50.382 | 49.618 | 0.400 | 1.609 |
| 2020-Q4 | 50.564 | 49.436 | 50.382 | 49.618 | 1.004 | 0.996 |
| Inactivity |  |  |  |  |  |  |
| T | $l_{i_{g, f}}$ | $l_{i_{g, m}}$ | $s_{i_{g, w}}$ | $s_{i_{g, m}}$ | $Q L_{w}$ | $Q L_{m}$ |
| 2020-Q2 | 35.214 | 64.786 | 60.603 | 39.397 | 0.581 | 1.644 |
| 2020-Q4 | 51.459 | 48.541 | 60.603 | 39.397 | 0.849 | 1.232 |

Table 1: Quarter loss in employment, structural unemployment and inactivity by gender considering both the second and the last quarter of 2020 as trough of the recession.

According to our estimates, the Covid-19 crisis has an impact that is more than proportional for female employment, while less than proportional for men considering both recession periods. The job losses for women account for $42.9 \%$ when considering only the first two quarters of 2020 and $57.28 \%$ when considering all quarters of 2020 as recession duration. Being actual employment at the prior-recession peak $42.16 \%$ for women, the QL ratio is slightly above 1 in the first case (1.018) and quite consistently above 1 (1.36) in the second case, signalling that with respect to the share of employment before the recession the
job losses have been disproportionate for women. Results for structural unemployment are less straightforward: both female and male structural unemployment decrease and are lower than the trend values (see Figure 15b and Figure B. 4 in the Appendix). The decrease has a higher impact on men considering the 2020-Q2 recession period, while both indicators are close to unity considering the 2020-Q4 recession period. Considering the entire time span, it looks like that there are no relevant gender differences in the impact of the Covid-19 crisis on structural unemployment being in both cases close to the unitary values. The decrease in structural unemployment is mirrored in the increase in inactivity, especially for men. The indicators suggest that the impact is more than proportional for men, especially considering the first time span (see Figure 15c), while it is less than proportional for women. This is given by the high hysteresis in female inactivity: considering the second estimate, the loss function is higher for women, as the share at the prior-recession peak is $60.60 \%$ for women vis-à-vis $39.40 \%$ for men. The Hodrick-Prescott filter confirms the results (Table C. 6 in the Appendix).

Overall, considering (i) the disproportionate effect for female employment given the low pre-recession employment rate, (ii) the proportionate effect on women in inactivity given the high pre-recession inactivity rates, and (iii) similar losses in structural unemployment, the influence of the past and persistent gender asymmetries on the impact of the Covid-19 crisis actually represent an amplifier of the She-recession. The more than proportionate effect on male inactivity seems to support the evidence of high increase in male inactivity observed during the Covid-19 crisis (see Section 5.2).

### 5.3.1 By region, education and sectors

Now, we present some evidence of the impact on female employment by region, by education level and sectors. Table 2 shows the job quarter loss for women in the North (n), Centre (c) and South (s) of Italy. Women from the South have suffered an impact more than proportional with respect to employed women in other regions, considering both time spans, during the first quarters of 2020 the impact has been more than proportional for women from the centre of Italy as well. Figure 16 shows the quarter-loss areas.


Figure 16: Regional female employment data filtered up to $2020-\mathrm{Q} 4$ quarter loss in grey

| Regional female employment |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| T | $l_{\epsilon_{r, n}}$ | $l_{\epsilon_{r, c}}$ | $l_{\epsilon_{r, s}}$ | $s_{\epsilon_{r, n}}$ | $s_{\epsilon_{r, c}}$ | $s_{\epsilon_{r, s}}$ | $Q L_{n}$ | $Q L_{c}$ | $Q L_{s}$ |
| 2020-Q2 | 12.21 | 24.73 | 63.06 | 54.69 | 22.24 | 23.07 | 0.22 | 1.11 | 2.73 |
| 2020-Q4 | 39.04 | 17.93 | 43.03 | 54.69 | 22.24 | 23.07 | 0.71 | 0.81 | 1.87 |

Table 2: Quarter loss for female employment in different macro regions of Italy

Table 3 shows the quarter loss in female employment with respect to education.
Considering only the first two quarters of 2020 as recessionary periods, the impact is more than proportional for women with primary and lower secondary education, with a QL indicator much higher than 1 (11.24 and 3.07 respectively). Women with tertiary education seem to gain jobs instead. On the other hand, the quarter loss is more than proportional for women with tertiary education considering the entire time span, as for women with primary education. Results are shown in Figure 17 as well. The Hodrick-Prescott filter confirms the results (Table C. 7 and C. 8 in the Appendix) except for education patterns, showing the
different sensitivity of the filters (see Figure B. 6 in the Appendix). This disproportionate result for women with tertiary education derives from the education structure within sectors. Despite women with higher level of education are ex-ante expected to be more resilient to an event crisis, we suggest that the pandemic as hit on knowledge workers with tertiary education, as translators, tourist guides and consultants, being self-employed workers not covered by the firing restrictions applied by the Italian government. This consideration is supported by the disproportionate effect detected in education, human health and social work activities in the sectoral analysis. Table 5 in fact shows that education, health and social activities account for $51 \%$ of losses and the impact is disproportionate for women employed in such sector being the actual share before recession $31 \%$. Precarious contracts and childcare channel may explain the disproportionate effect.

Female employment by education

| T | $l_{\epsilon_{e, p}}$ | $l_{\epsilon_{e, l s}}$ | $l_{\epsilon_{e, u s}}$ | $l_{\epsilon_{e, t}}$ | $s_{\epsilon_{e, p}}$ | $s_{\epsilon_{e, l s}}$ | $s_{\epsilon_{e, u s}}$ | $s_{\epsilon_{e, t}}$ | $Q L_{p}$ | $Q L_{l s}$ | $Q L_{u s}$ | $Q L_{t}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2020-Q2 | 22.87 | 65.48 | 19.47 | -7.82 | 2.03 | 21.36 | 46.20 | 30.40 | 11.24 | 3.07 | 0.42 | -0.26 |
| 2020-Q4 | 9.24 | 11.27 | 37.66 | 41.83 | 2.03 | 21.36 | 46.20 | 30.40 | 4.54 | 0.53 | 0.82 | 1.38 |

Table 3: Quarter loss in female employment by education level

Table 4 and Figure 18 show the quarter losses in female employment in industries, constructions, agriculture and services. The impact has been more than proportional only for women employed in services, in which $84.4 \%$ of female workers are employed and accounting for a job loss share of $113 \%$, while women in constructions and industries have gained jobs instead (negative values). Results from Hodrick-Prescott filtered data are provided in Table C. 9 in the Appendix, according to which also women employed in the agriculture sector were subject to a disproportionate impact looking at only the first two quarters of 2020. The indicator is 1.05 considering the entire period. Considering the importance of the sectoral channel, we deepen the analysis looking at female employment in 9 different service sub-sectors.

According to Table 5, the quarter losses computed on the entire time span (recession trough 2020-Q4) are above unity for transportation and storage, accommodation and food activities, education, human health and social work activities, the latter being the highest sector of employment for women. At the opposite, information and communication, finance


Figure 17: Female employment by education level, data filtered up to 2020-Q4 quarter loss in grey
and insurance activities and public administration show instead an increase in employment with respect to predicted data, being such jobs possible to be performed remotely. Estimates with the Hodrick-Prescott filter (Table C. 10 in the Appendix) confirm the disproportionate impact for transportation and storage, accommodation and food but the indicator is above unity also for other services and administrative and support service activities. The indicator is instead below the unity for education, health and social work, despite it accounts for an important loss during the third quarter of 2020. As for tertiary education, this is the result of the different sensitivity of the two filters (see Figure B. 7 in the Appendix), the HP being more linear.

| Female employment by macro-sectors |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| T | $l_{\epsilon_{s, a}}$ | $l_{\epsilon_{s, c}}$ | $l_{\epsilon_{s, i}}$ | $l_{\epsilon_{s, s}}$ | $s_{\epsilon_{s, a}}$ | $s_{\epsilon_{s, c}}$ | $s_{\epsilon_{s, i}}$ | $s_{\epsilon_{s, s}}$ | $Q L_{a}$ | $Q L_{c}$ | $Q L_{i}$ | $Q L_{s}$ |
| 2020-Q2 | 0.59 | -10.58 | -3.22 | 113.21 | 2.62 | 1.01 | 12.33 | 84.04 | 0.23 | -10.45 | -0.26 | 1.35 |
| 2020-Q4 | 0.65 | -5.35 | 2.23 | 102.47 | 2.62 | 1.01 | 12.33 | 84.04 | 0.25 | -5.29 | 0.18 | 1.22 |

Table 4: Quarter loss in female employment by macro-sectors

According to Table 5, the quarter losses computed on the entire time span (recession


Figure 18: Female employment by macro-sectors, data filtered up to 2020-Q4, quarter loss in grey.
trough 2020-Q4) are above unity for transportation and storage, accommodation and food activities, education, human health and social work activities, the latter being the highest sector of employment for women. At the opposite, information and communication, finance and insurance activities and public administration show instead an increase in employment with respect to predicted data, being such jobs possible to be performed remotely. Estimates with the Hodrick-Prescott filter (Table C. 10 in the Appendix) confirm the disproportionate impact for transportation and storage, accommodation and food but the indicator is above unity also for other services and administrative and support service activities. The indicator is instead below the unity for education, health and social work, despite it accounts for an important loss during the third quarter of 2020. As for tertiary education, this is the result of the different sensitivity of the two filters (see Figure B. 7 in the Appendix), the HP being more linear.

### 5.3.2 Distribution of disproportionate effects

This sub-section has the purpose to show the distribution of our measure around the unity in order to check the incidence of proportionality vs disproportionality events. We provide for the distribution of the QL indicator, checking how often it takes values above or below

| Sector | $l_{\epsilon_{\text {serv }}}$ | $s_{\epsilon_{\text {serv }}}$ | $Q L_{\text {serv }}$ |
| :--- | :--- | :--- | :--- |
| Wholesale and retail trade and repair of motor vehicles and motorcycles | 12.3 | 16.94 | 0.73 |
| Transportation and storage | 5.57 | 2.9 | 1.92 |
| Accommodation and food activities | 10.72 | 8.38 | 1.28 |
| Information and communication | -1.73 | 2.05 | -0.85 |
| Financial and insurance activities | -3.63 | 3.53 | -1.03 |
| Public administration and defence; compulsory social secutiry | -0.05 | 4.98 | -0.01 |
| Administrative and support service activities | 13.79 | 15.57 | 0.89 |
| Other services | 11.37 | 14.6 | 0.78 |
| Education, Human health and social work activities | 51.68 | 31.07 | 1.66 |

Table 5: Quarter loss in female employment in service sub-sectors
the unity in order to have a complete assessment of the range of its dispersion among different tests of disproportionality. Figure 20 shows the distribution of the indicator when considering the second quarter of 2020 as last quarter of the recession period ( $\mathrm{T}=2020-\mathrm{Q} 2$, on the left) vis-à-vis the last quarter of 2020 ( $\mathrm{T}=2020-\mathrm{Q} 4$ on the right), and the distribution of all indicators as a third graph. Not surprisingly, the longest recession reports far larger extreme values than the short one. The distribution of all indicators is concentrated between 0.8 and 1.3. Disproportional values (larger than unity) are present but are also balanced by less than proportional ones, considering that we are including the all range of values of the indicator. Outliers belongs particularly to the sectoral distribution.

## 6 Conclusions and policy implications

Since the Covid-19 economic crisis has been prompted by an unprecedented health emergency, it has impacted on the economic system in a completely different manner when compared to standard downturns: in particular, for the first time a stronger impact on female employment has been recorded, while during past recessions, male employment has shown much higher cyclicality (Man-recession). For this reason, the Covid-19 crisis has been defined as a She-recession, because of (i) the industry channel, for which the sectors hit the hardest by social distancing measures are characterised by high shares of female


Figure 19: Female employment by service sectors, data filtered up to 2020-Q4, quarter loss in grey


Figure 20: Quarter loss indicators distribution and density function above and below the unity
employment, and (ii) the childcare channel, i.e., school closures have increased the childcare burden especially on women because of gender norms, triggering transition to inactivity.

However, considering the impact of the recession in an isolated manner, without accounting for pre-pandemic structural asymmetries in gender, geographical distribution and education, risks to undervalue the root causes of the recorded patterns during the pandemic. While the literature has widely documented gender imbalances in labour market outcomes, as the gender-pay gap and the lower female participation rate, matters of feminization of the labour markets have been less considered (Cetrulo et al. (2023)). In addition, although internationally a wide coverage has been devoted to document the She-recession, less evidence is available for Italy, and particularly with a perspective drawing upon long-term memory processes and hysteresis as fuelling amplifiers of "temporary" shocks.

This paper aims at measuring and explaining the gender differences in the impact of the Covid-19 crisis on the Italian labour market from a macroeconomic perspective by assessing the depth, duration and diffusion of the recession. In particular, we trace back to pre-existing and persistent gender asymmetries the roots of the She-recession. Our main findings suggest that hysteresis in gender asymmetries amplifies and affects the She-recession manifestation of the Covid-19 crisis, given the disproportionate impact on female employment and proportionate effect on female inactivity. In line with hysteresis in labour markets, women from the South and with lower education levels suffer an impact more than proportionate with respect to higher educated women, despite the impact on graduated women is more than proportional when considering all the quarters of 2020 as recession period. On the one hand, the impact on low educated women can be explained by the industry channel, since they are mostly employed in low value added activities mainly hit by the pandemic restrictions. A similar reasoning applies for women working in the South, being their activities concentrated in sectors subject to closures. For women with tertiary education, the children channel could be an explanation as the high share of knowledge self-employed workers not covered by the firing restrictions. Evidence from sectoral analysis suggests a disproportionate impact on women working in the service sector and in particular in transportation and storage, accommodation and food services and but even in education, human health and social work activities. Such result may explain and be explained by the disproportionate effect on women with tertiary education and the childcare channel. This is out of the scope of this paper, but worthy for further investigation.

More broadly, the results of our analysis indicate, firstly, the structural vulnerability of the female segment in the labour market. Secondly, vulnerabilities are not commonly widespread, and inside the female component, women from the South and employed in low remunerated service sectors have been the most exposed to job losses, calling for a multidimensional perspective of exposure to risks (Cetrulo et al. (2022)). In addition, even the broadly defined care sector, including health, education and social care works has been disproportionately hit by the pandemic shock, showing that the so called essentiality narrative of these jobs has not impeded the adverse effects of the pandemic. Thirdly, although the disproportionate impact on women has been documented word-widely, country effects have been different, and in general, whenever more comprehensive welfare state regimes were in place, the She-recessionary effects have been lower. At the opposite, whenever informality and self-employment were higher, the effects of the She-recession have been more severe (e.g., the case of the South of Italy). Granted that the welfare state and care regimes interact, the She-recession has been proved to be deeply related to the ex-ante hysteresis and pre-crisis conditions in the labour market, a phenomenon defined as long-term feminization.

Our results call therefore for policy actions that are meant to overturn long-term feminization, namely, both intra-household and extra-household unequal status of women, mapping into gender asymmetries and discrimination. Beyond the simplistic factual consideration of gender heterogeneity and gender-wage gap, there are deeper institutional, cultural and social norms which reproduce gender imbalances and subordination to unequal societal power (Folbre (2021)), call it "patriarchy cum neo-liberalism". Still nowadays, gender asymmetries are to a larger extent considered the end result of individual choices, while structurally embedded societal factors, leading to the trap into feminization, have not put under the spotlight with sufficient attention. Given that women are subject to discrimination in their hiring processes and career advancement, in accepting involuntary part-time contracts and in the decision to enter or not in the labour force because of family culture and constraints, societies should question the acceptability of such status quo and the rootcauses that allow gender discrimination to keep reproducing. In addition, female subjects are even more exposed to adverse effects whenever they are not in a couple and with children, therefore whenever they lack the male counterpart as a presumed element of individual "security". Labour market outcomes and household status intersect with other spheres, such as access to healthcare and childcare, reproducing the so-called intersectionality and
multi-dimensionality of vulnerability.
So far, policy interventions have been very timid, if not completely absent, in overturning structural gender imbalances ${ }^{18}$ : while the discourse has usually focused on career upgrading, gender quota, STEM education and female entrepreneurship, no advancement is put forward toward substantive processes of women empowerment, granting higher spaces of decision-making autonomy, equal rights and, in that, overturning occupational segregation. At the same time, no economic policy proposal has envisaged an industrial policy for the care sector, able to grant both equal access to basic rights, as education and health, and good and stable employment conditions for women (Cresti and Virgillito (2022)). An industrial policy for the care sector would allow to tackle at the same time labour market asymmetries for women and access to basic needs for all, requesting for the State to be a good employer. While policy actions have been mostly lacking, a Care manifesto and the notion of a care economy have been emerging in the last years, intersecting the political and academic debates (Chatzidakis et al. (2020). In essence, the care economy highlights interdependence among humans, humans and nature, and the need to overturn crystallised class structures and power asymmetries in contemporary societies. Affirming interdependence, the Care manifesto calls for a new regime of care in which unpaid working time and low value jobs are deeply revalued and remunerated. However, such policy actions, more than simply fixing market failures, would entail a collective reconsideration of the nature of the current structure of social and economic relations, what Folbre (2021) defines "bargaining for changes".$^{19}$

Occupational-level, cross-country comparative analyses and measurement of long-lasting She-recessionary effects are further avenues of research together with a deeper accounting of the impact of feminization, largely intended both as a process occurring extra-household in the labour market, but even intra-household, with reference to the gendered division of unpaid labour and effects of gender norms into participation to labour markets.

[^11]
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## Appendix

## A The Hodrick-Prescott filter

Assume we have a time series

$$
\begin{equation*}
x_{t}=y_{t}+\bar{x}_{t} \tag{9}
\end{equation*}
$$

Spectral analysis provides as instrument to obtain the component $y_{t}$, an ideal bandpass filter:

$$
\begin{equation*}
y_{t}=B(L) x_{t} \tag{10}
\end{equation*}
$$

where $B(L)=\sum_{j=-\infty}^{\infty} B_{j} L^{j}$ is the filter, $L$ is the lag operator and $B_{j}=\frac{\sin (j b)-\sin (j a)}{\pi j}$, $B_{0}=\frac{b-a}{\pi}$ are the weights. However, the the ideal bandpass filter can be applied only if data are infinite. The implementation of the Christian-Fitzgerald and the Hodrick-Prescott (CF and HP from now on) starts from the necessity to have an instrument able to make this decomposition when data are finite. As a result, these filters apply an approximation of the ideal one (Fitzgerald and Christiano (1999)).

The HP filter has the aim to explain fluctuations of aggregate macroeconomic variables over the business cycle, from the long run path of growth (Hodrick and Prescott (1997)).

Time series are represented as the sum of a growth and a cyclical component

$$
\begin{equation*}
x_{t}=g_{t}+c_{t} \quad t=1, \ldots, T \tag{11}
\end{equation*}
$$

where growth changes smoothly and the cycle component is defined as the deviation form growth path. The aim is to minimise the smoothness of growth, that is:

$$
\begin{equation*}
\min _{\left\{g_{t}\right\}_{t=-1}^{T}}\left\{\sum_{t=1}^{T} c_{t}^{2}+\lambda \sum_{t=1}^{T}\left[\left(g_{t}-g_{t-1}\right)-\left(g_{t-1}-g_{t-2}\right)\right]^{2}\right\} \tag{12}
\end{equation*}
$$

where $\lambda$ is a penalty parameter, that is a positive number that the higher the variations in growth component the higher the penalty. Assuming the cycle component and the second difference of the growth component with zero mean and constant variances $\sigma_{1}^{2}$ and $\sigma_{2}^{2}$, the penalty parameter is defined as $\sqrt{\lambda}=\frac{\sigma_{1}}{\sigma_{2}}$. For quarterly data, Hodrick and Prescott consider a 5 percent cyclical component and $\frac{1}{8}$ of one percent change in growth rate in a quarter as moderately large, hence $\sqrt{\lambda}=\frac{5}{d \frac{1}{8}=40}$, that is $\lambda=1,600$.

## B Figures



Figure B.1: Female employment rate, age 15 and over, comparison between EU-2020 ( 27 members) average and Italy. Data Source: European Labour Force Survey, Eurostat


Figure B.2: Cyclical and trend variation of the Italian GDP by quarters from 1996 to 2020


Figure B.3: Filtered employment data up to the quarter 2020-Q2 by gender by the ChristianoFitzgerald filters, quarter loss in grey.


Figure B.4: Filtered structural unemployment data up to the quarter 2020-Q2 by gender by Christiano-Fitzgerald filter, quarter loss in grey.


Figure B.5: Filtered inactivity data up to the quarter 2020-Q2 by gender by the ChristianoFitzgerald filter, quarter loss in grey


Figure B.6: Female employment by education level, filtered data by Hodrick-Prescott filter up to 2020-Q4, quarter loss in grey. Data Source: Labour Force Survey, ISTAT


Figure B.7: Female employment in service sub-sectors filtered data by Hodrick-Prescott filter up to 2020-Q4, quarter loss in grey

## C Tables

Table C.1: Descriptive statistics of the total, male and female employment, structural unemployment and inactivity, quarterly data, age 15 and over. Data Source: Labour Force Survey, ISTAT

|  | Employment |  |  |
| :--- | :---: | :---: | :---: |
| Statistic | Female | Male | Total |
|  |  |  |  |
| t: 1993-Q1/2020-Q1 | $\mathrm{t}=112$ | $\mathrm{t}=112$ | $\mathrm{t}=112$ |
| Mean | $8,804,935$ | $13,377,585$ | $22,182,520$ |
| Median | $9,065,017$ | $13,353,464$ | $22,405,984$ |
| St. Dev. | $736,713.400$ | $253,506.200$ | $802,131.100$ |
| Min | $7,398,284$ | $12,766,801$ | $20,522,766$ |
| Max | $9,997,537$ | $13,939,058$ | $23,553,667$ |


|  | Structural unemployment |  |  |
| :--- | :---: | :---: | :---: |
| Statistic | Female | Male | Total |
|  |  |  |  |
| t: 1993-Q1/2020-Q1 | $\mathrm{t}=112$ | $\mathrm{t}=112$ | $\mathrm{t}=112$ |
| Mean | $604,888.900$ | $602,806.400$ | $1,207,695$ |
| Median | 608,514 | 550,275 | $1,161,076$ |
| St. Dev. | $144,883.100$ | $209,565.500$ | $349,337.600$ |
| Min | 332,883 | 290,625 | 623,508 |
| Max | $1,022,043$ | $1,093,141$ | $2,062,632$ |
|  | Femactivity |  |  |
| Statistic |  |  |  |
|  | $\mathrm{t}=112$ | $\mathrm{t}=112$ | $\mathrm{t}=112$ |
| $\mathrm{t}:$ 1993-Q1/2020-Q1 | $15,961,169$ | $9,465,930$ | $25,427,099$ |
| Mean | $15,960,906$ | $9,440,054$ | $25,483,066$ |
| Median | $275,743.700$ | $705,876.000$ | $898,781.700$ |
| St. Dev. | $15,417,757$ | $8,186,644$ | $24,095,277$ |
| Min | $16,596,764$ | $10,891,029$ | $27,410,160$ |
| Max |  |  |  |

Table C.2: Descriptive statistics of female regional employment, quarterly data, age 15 and over. Data Source: Labour Force Survey, ISTAT

| Female regional employment |  |  |  |
| :--- | :---: | :---: | :---: |
| Statistic | North | Centre | South |
| t: 2000-Q1/2020-Q4 | $\mathrm{t}=84$ | $\mathrm{t}=84$ | $\mathrm{t}=84$ |
| Mean | $4,986,625$ | $2,004,683$ | $2,177,057$ |
| Median | $5,027,856$ | $2,020,519$ | $2,169,161$ |
| St. Dev. | $219,108.200$ | $152,710.800$ | $70,179.390$ |
| Min | $4,433,069$ | $1,633,726$ | $1,959,192$ |
| Max | $5,391,286$ | $2,252,968$ | $2,353,283$ |

Table C.3: Descriptive statistics of female employment by education level, seasonally adjusted, quarterly data, age 15 and over. Data source: Labour Force Survey, ISTAT

| Female employment by education |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Statistic | Primary | Lower Secondary | Upper Secondary | Tertiary |
|  |  |  |  |  |
| $\mathrm{t}: 2000-\mathrm{Q} 1 / 2020-\mathrm{Q} 4$ | $\mathrm{t}=84$ | $\mathrm{t}=84$ | $\mathrm{t}=84$ | $\mathrm{t}=84$ |
| Mean | $449,410.400$ | $2,250,118$ | $4,367,995$ | $2,100,842$ |
| Median | 388,812 | $2,244,788$ | $4,453,724$ | $2,079,064$ |
| St. Dev. | $217,706.200$ | $132,796.800$ | $221,956.800$ | 557,931 |
| Min | 164,703 | $1,930,380$ | $3,704,025$ | $1,175,015$ |
| Max | 877,990 | $2,551,847$ | $4,642,923$ | $3,106,937$ |

Table C.4: Descriptive statistics for female employment in macro-sectors, seasonally adjusted, quarterly data, age 15 and over. Data source: Labour Force Survey, ISTAT

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Statistic | Agriculture | Industry (no constructions) | Constructions | Services |
| $\mathrm{t}: 2008-\mathrm{Q} 1 / 2020-\mathrm{Q} 4$ | $\mathrm{t}=52$ | $\mathrm{t}=52$ | $\mathrm{t}=52$ | $\mathrm{t}=52$ |
| Mean | $238,012.700$ | $1,194,168.000$ | $102,475.900$ | $7,900,942.000$ |
| St. Dev. | $21,970.100$ | $59,253.030$ | $12,857.300$ | $270,898.500$ |
| Min | 192,242 | $1,121,008$ | 82,309 | $7,461,804$ |
| Max | 285,696 | $1,393,379$ | 138,650 | $8,482,943$ |


| Employment |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| T | $l_{\epsilon_{g, w}}$ | $l_{\epsilon_{g, m}}$ | $s_{\epsilon_{g, w}}$ | $s_{\epsilon_{g, m}}$ | $Q L_{w}$ | $Q L_{m}$ |  |
| $2020-\mathrm{Q} 2$ | 47.878 | 52.122 | 42.157 | 57.843 | 1.136 | 0.901 |  |
| $2020-\mathrm{Q} 4$ | 50.723 | 49.277 | 42.157 | 57.843 | 1.203 | 0.852 |  |

Structural unemployment

| T | $l_{u_{g, w}}$ | $l_{u_{g, m}}$ | $s_{u_{g, w}}$ | $s_{u_{g, m}}$ | $Q L_{w}$ | $Q L_{m}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2020-Q2 | 52.880 | 47.120 | 50.382 | 49.618 | 1.050 | 0.950 |
| 2020-Q4 | 53.053 | 46.947 | 50.382 | 49.618 | 1.053 | 0.946 |
| Inactivity |  |  |  |  |  |  |
| T | $l_{i_{g, w}}$ | $l_{i_{g, m}}$ | $s_{i_{g, w}}$ | $s_{i_{g, m}}$ | $Q L_{w}$ | $Q L_{m}$ |
| $2020-\mathrm{Q} 2$ | 50.840 | 49.160 | 60.603 | 39.397 | 0.839 | 1.248 |
| $2020-\mathrm{Q} 4$ | 53.649 | 46.351 | 60.603 | 39.397 | 0.885 | 1.177 |

Table C.6: Quarter loss with respect to the Hodrick-Prescott filter in employment, structural unemployment and inactivity

| Statistic | t:2008Q1/2020Q4 | Mean | St. Dev. | Min | Max |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Wholesale and retail trade, repair of motorvehicles and motorcycles | $\mathrm{t}=52$ | 1,350,212.000 | 38,889.180 | 1,263,695 | 1,437,968 |
| Transportation and Storage | $\mathrm{t}=52$ | 215,664.700 | 14,868.440 | 186,070 | 247,524 |
| Accommodation and food services | $\mathrm{t}=52$ | 657,866.900 | 69,819.430 | 537,172 | 808,972 |
| Information and Communication | $\mathrm{t}=52$ | 173,919.700 | 12,447.250 | 142,509 | 205,232 |
| Financial and insurance activities | $\mathrm{t}=52$ | 282,109.400 | 12,803.440 | 251,640 | 314,483 |
| Administration and support services | $\mathrm{t}=52$ | 1,204,422.000 | 65,243.330 | 1,083,271 | 1,346,365 |
| Public administration and defense | $\mathrm{t}=52$ | 451,678.900 | 29,169.130 | 403,661 | 504,997 |
| Education, human health and social work | $\mathrm{t}=52$ | 2,406,468.000 | 118,035.900 | 2,209,399 | 2,700,402 |
| Other services | $\mathrm{t}=52$ | 1,158,601.000 | 96,157.810 | 875,814 | 1,278,327 | source: Labour Force Survey, ISTAT


| Regional female employment |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| T | $l_{\epsilon_{r, n}}$ | $l_{\epsilon_{r, c}}$ | $l_{\epsilon_{r, s}}$ | $s_{\epsilon_{r, n}}$ | $s_{\epsilon_{r, c}}$ | $s_{\epsilon_{r, s}}$ | $Q L_{n}$ | $Q L_{c}$ | $Q L_{s}$ |
| 2020-Q2 | 40.696 | 22.715 | 36.58 | 54.691 | 22.239 | 23.070 | 0.744 | 1.021 | 1.586 |
| 2020-Q4 | 47.948 | 23.169 | 28.88 | 54.691 | 22.239 | 23.070 | 0.877 | 1.042 | 1.252 |

Table C.7: Quarter loss by Hodrick-Prescott filter for female employment in different macro regions of Italy

| Female employment by education |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| T | $l_{\epsilon_{e, p}}$ | $l_{\epsilon_{e, l s}}$ | $l_{\epsilon_{e, u s}}$ | $l_{\epsilon_{e, t}}$ | $s_{\epsilon_{e, p}}$ | $S_{\epsilon_{e, l_{s}}}$ | $s_{\epsilon_{e, u s}}$ | $s_{\epsilon_{e, t}}$ | $Q L_{p}$ | $Q L_{l s}$ | $Q L_{u s}$ | $Q L_{t}$ |
| 2020-Q2 | 7.639 | 33.874 | 33.360 | 25.127 | 2.035 | 21.363 | 46.197 | 30.405 | 3.754 | 1.586 | 0.722 | 0.826 |
| 2020-Q4 | 3.765 | 31.349 | 31.365 | 33.521 | 2.035 | 21.363 | 46.197 | 30.405 | 1.850 | 1.467 | 0.679 | 1.102 |

Table C.8: Quarter loss with respect to the Hodrick-Prescott filter in female employment by education level

| Female employment by macro-sectors |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| T | $l_{\epsilon_{s, a}}$ | $l_{\epsilon_{s, c}}$ | $l_{\epsilon_{s, i}}$ | $l_{\epsilon_{s, s}}$ | $s_{\epsilon_{s, a}}$ | $s_{\epsilon_{s, c}}$ | $s_{\epsilon_{s, i}}$ | $s_{\epsilon_{s, s}}$ | $Q L_{a}$ | $Q L_{c}$ | $Q L_{i}$ | $Q L_{s}$ |
| 2020-Q2 | 13.89 | -12.24 | 0.44 | 97.91 | 1.01 | 12.33 | 2.62 | 84.04 | 5.30 | -12.09 | 0.04 | 1.17 |
| 2020-Q4 | 2.76 | -4.85 | 6.19 | 95.89 | 1.01 | 12.33 | 2.62 | 84.04 | 1.05 | -4.79 | 0.50 | 1.14 |

Table C.9: Quarter loss by Hodrick Prescott filter for female employment by macro-sectors

| Sector | $l_{\epsilon_{\text {serv }}}$ | $s_{\epsilon_{\text {serv }}}$ | $Q L_{\text {serv }}$ |
| :--- | :--- | :--- | :--- |
| Wholesale and retail trade and repair of motor vehicles and motorcycles | 9.27 | 16.94 | 0.55 |
| Transportation and storage | 3.28 | 2.9 | 1.13 |
| Accommodation and food activities | 36.7 | 8.38 | 4.38 |
| Information and communication | -1.41 | 2.05 | -0.69 |
| Financial and insurance activities | -1.81 | 3.53 | -0.51 |
| Public administration and defence; compulsory social secutiry | 2.01 | 4.98 | 0.4 |
| Administrative and support service activities | 18.36 | 15.57 | 1.18 |
| Other services | 25.23 | 14.6 | 1.73 |
| Education, Human health and social work activities | 8.38 | 31.07 | 0.27 |

Table C.10: Quarter loss by Hodrick Prescott filter for female employment by service subsectors filtered up to 2020-Q4


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[^1]:    ${ }^{1}$ Total employment has decreased by 101 thousand units in December 2020 with respect to November 2020, of which 99thousand units of female employment (cyclical variation in Table 1) ISTAT.pdf
    ${ }^{2}$ In the poverty literature, the term feminization relates to the concept of feminization of poverty (Pearce (1978) as cited byMc Lanahan et al. (1989)) for which the exposure to poverty risk for women is higher than for men because of the change in family structures, occupational segregation on the labour market and welfare programs Peterson (1987)). While the implications for poverty risk are not addressed in our analysis, in this paper, we refer to feminization of the labour market. The concept intends that the increasing female participation rate in the labour market has been accompanied by the servitization and flexibilization of employment relations since the 1990s. In that, feminization implies occupational segregation and precariousness disproportionally affecting women (Manicardi, 2023 this issue; Rubery (2015); Betti (2016); Cetrulo et al. (2023)). Labour market feminization has been however also ignited by gender norms implying low participation rates and part-time working arrangements. In that, we uncover both extra-household and intra-household relations leading to feminization.

[^2]:    ${ }^{3}$ B. 1 in the Appendix

[^3]:    ${ }^{4}$ The initial time and age differ in the discussed trends because of data availability for each category. 2020 is always the last period since we analyse the effect of the Covid-19 crisis during its first year. Employment by sectors and professional status are all age 15 and over to harmonize with availability of data.
    ${ }^{5}$ The NACE specification is the corresponding English translation of the sectors classified under the service macro-sector by ISTAT

[^4]:    ${ }^{6}$ We look at structural unemployment as a variable characterised by strong hysteresis since the focus of the paper is on the long term patterns of feminization of the labour market. We identify hysteresis as one of the main roots of the Covid-19 crisis unfolding as a She-recession and in that looking at structural rather than cyclical unemployment. The Labour Force Survey distinguishes for duration of unemployment shorter or longer than 12 months. Further details are provided in Subsection 5.1. The age cohort differs from inactivity because of data availability.

[^5]:    ${ }^{7}$ We choose the cohort of age of 15 and over since it is the one that is common to all three variables, since age 15-64 in not available for structural unemployment. By the same token, Istat do not provide data for long-term unemployment from 1977 as for inactivity and employment, hence we use data from 1993

[^6]:    ${ }^{8}$ see Table C. 1
    ${ }^{9}$ For instance, in 2019 before the pandemic, female employment in Italy was $53.8 \%$ vis-à-vis $67.5 \%$ as average in the European Union, 27 members, Eurostat (Figure B. 1 in the Appendix)
    ${ }^{10}$ age 15-64, inactive workers in March 2019 were $13,016,046$ and 14,223,474 in March 2020, data source: Labour Force Survey, ISTAT. $10 \%$ of trend variation of inactive workers during the second quarter of 2020 , corresponding to 1 million and 310 thousands units more with respect to 2019https://www.istat.it/it/ files/2020/09/Mercato-del-lavoro-II-trim_2020.pdf

[^7]:    ${ }^{11}$ including the activities of membership organisations, the repair of computers and personal and household goods and a variety of personal service activities, activities of households as employers, undifferentiated goods and services producing activities of households for own use

[^8]:    ${ }^{12}$ NBER Business Cycle Dating Committee guidelines

[^9]:    ${ }^{13}$ According to Christiano and Fitzgerald (2003), for what concerns the analysis of business cycles with

[^10]:    ${ }^{14}$ Islands are included as southern regions.
    ${ }^{15} 2020-\mathrm{Q} 4$, if considering the trend variation of GDP to define the recession period (see step 1)
    ${ }^{16} 2020-\mathrm{Q} 2$, if considering the cyclical variation to define the recession period, (see step 1 ).
    ${ }^{17}$ Fazzari and Needler (2021) define their measure as job monthly loss.

[^11]:    ${ }^{18}$ See https://feps-europe.eu/wp-content/uploads/2022/12/RECOVERY-WATCH-Feminist-Care-PP-2. pdf
    ${ }^{19}$ See Folbre 2021, "Gender inequality and bargaining in the U.S. labor market", Economic Policy Institute, https://www.epi.org/unequalpower/publications/gender-and-bargaining-in-the-u-s-labor-market/

